

PnS: Assignment 1

g1) $X \sim \text{Poi}(\theta)$

$$\frac{X - \theta}{\sqrt{\theta}} \sim N(0, 1)$$

$$X = 200$$

$$\therefore \theta = X \pm 1.96(\sqrt{\theta/n})$$

we also have

$$\begin{aligned} \text{var}\left[\frac{(n-1)s^2}{\sigma^2}\right] &= 2(n-1) \Rightarrow \text{var}[s^2] = \frac{\sigma^4}{(n-1)^2} \times 2(n-1) \\ &= \frac{2\sigma^4}{(n-1)} \end{aligned}$$

Q2) $X_i = 1, 2, \dots, n$

mean = μ

var = σ^2

$\bar{X} = \sum_{i=1}^n X_i = \text{sample mean}$

$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \text{sample variance}$

(i) S.T mean of $\bar{X} = \mu$ & var = $\frac{\sigma^2}{n}$

→ Since it is a sequence of independent, i.d.r.v.s with mean μ and variance σ^2 then the distribution of $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$ approaches standard normal distribution, $N(0,1)$

since $n \rightarrow \infty$

$\therefore E(\bar{X}) = \frac{E(\sum X_i)}{n} = \frac{n\mu}{n} = \mu$

$\therefore V(\bar{X}) = \frac{\text{Var}(\sum X_i)}{n^2} = \frac{1}{n^2} (n\sigma^2) = \frac{\sigma^2}{n}$

(ii) $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{n-1} \quad \therefore E(S^2) = \frac{(n-1)\sigma^2}{(n-1)} = \sigma^2$

$\therefore \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$ are approximately distributed as $N(0,1)$

$\therefore \bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$

$\therefore \text{Mean} = \underline{\underline{\mu}}; \text{Var} = \underline{\underline{\frac{\sigma^2}{n}}} \quad \text{f } E(S^2) = \sigma^2$

(iii) S.T $\text{Var}(S^2) = \frac{2\sigma^4}{(n-1)}$

→ Using χ^2 result,

$E\left[\frac{(n-1)S^2}{\sigma^2}\right] = n-1 \Rightarrow E[S^2] = \frac{\sigma^2}{n-1} (n-1) = \sigma^2$ and

Given: $n=10$; $\sigma^2=100$

$$(iv) P(50 < S^2 < 150) = P\left(\frac{9 \times 50}{100} < \chi^2_9 < \frac{9 \times 150}{100}\right)$$

$$= P\left(\frac{9}{2} < \chi^2_9 < \frac{27}{2}\right)$$

$$= \underline{\underline{4.5, 13.5}}$$

Q3) Binomial, $n=4$ and p . $\therefore \text{Bin}(4, p)$ no. of policies = 180

$$\rightarrow (i) L(p) = \sum_{i=0}^4 \frac{[1-p]^4}{4!} \cdot [4p^i(1-p)^{4-i}]^4$$

$$= [(1-p)^4]^{86} \cdot [4p(1-p)^3]^{75} \cdot [6p^2(1-p)^2]^{16} \cdot [4p^3(1-p)]^2 \cdot p^4$$

$$= (1-p)^{344} \cdot p^{75} (1-p)^{225} \cdot p^{32} (1-p)^{32} \cdot p^6 (1-p)^2 \cdot p^4$$

$$= (1-p)^{603} \cdot p^{117}$$

$$\ln L(p) = 603 \ln(1-p) + 117 \ln p$$

$$\therefore \frac{d}{dp} \ln L(p) = \frac{-603}{1-p} + \frac{117}{p} = 0$$

$$\therefore 117 - 117p - 603p$$

$$\therefore q = 0.8375 \quad \therefore p = 0.1625$$

$$(ii) P(X=0) = (1-p)^4 = 0.492$$

$$P(X=1) = 4p(1-p)^3 = 0.3818$$

$$P(X=2) = 6p^2(1-p)^2 = 0.1111$$

$$P(X=3) = 4p^3(1-p) = 0.0144$$

$$P(X=4) = 0.0007$$

we have

O_i	86	75	16	2	1
E_i	88.56	68.72	19.98	2.59	0.126

By clubbing,

O_i	86	75	19
E_i	88.56	68.72	22.71

$$C_i = \frac{\sum (O_i - E_i)^2}{E_i} \quad \left| \begin{array}{c} 0 \\ 0.0757 \\ 0.5739 \\ 0.6061 \end{array} \right|$$

$$\sum C_i = 1.2557 \sim \chi^2_1 \quad \therefore 1.2557 < 3.841$$

$\therefore$ do not reject H_0 .

Q5) Since $U(a, b)$,

(i) Given that $E(X) = \frac{a+b}{2}$; $Var(X) = \frac{(b-a)^2}{12}$

show that $\hat{a} = \bar{x} - \sqrt{3}s$ where $\bar{x}$ = sample mean
 s = sample S.D.

$$\text{Having } a = 2\bar{x} - b$$

$$\therefore a = b - 2\sqrt{3}s$$

$$\therefore 2\bar{x} - b = b - 2\sqrt{3}s$$

$$\therefore b = \bar{x} + \sqrt{3}s$$

$$\therefore \underline{a = \bar{x} - \sqrt{3}s}$$

Hence proved.

(ii) $b = \bar{x} + \sqrt{3}s$

(iii) let sample = 2, 4, 6, 8, 10

$$\therefore \bar{x} = 6$$

$$s = 3.16$$

$$\therefore a = 6 - \sqrt{3}(3.16)$$

$$= \underline{0.527}$$

$$\therefore b = 6 + \sqrt{3}(3.16)$$

$$= \underline{11.47}$$

(iv) $U(a, b)$ $\therefore f(x) = \frac{1}{b}$

$$\therefore L = \frac{1}{b^n}$$

$$\log L = -n \log b$$

$$\therefore \frac{d}{dL} \log L = \frac{-n}{b} \quad \therefore \underline{b \rightarrow \alpha}$$

Q7)

(i) $\bar{x}_{\text{public}} = 30$

$s^2_{\text{public}} = 64.3125$

$\bar{x}_{\text{private}} = 35.056$

$s^2_{\text{private}} = 67.403$

(ii) we want $H_0: \mu_{\text{pub}} = \mu_{\text{pri}}$

$H_1: \mu_{\text{pub}} \neq \mu_{\text{pri}}$

$$\therefore \frac{(30 - 35.056)}{\sqrt{8.123} \sqrt{\frac{2}{9}}} = -1.407$$

$\therefore$ ~~do~~ do not reject H_0 .

(iii) Private hospitals do not result in higher claim size at 95% level.

Q9) (i) $t_k = \frac{N(0,1)}{\sqrt{\frac{t_k}{k}}}$; $\frac{\bar{x} - \mu}{s/\sqrt{n}} \sim t_{n-1}$

(ii) $\mu(t_k) = 0$; $\text{Var}(t_k) = \frac{k}{k-2}$

(iii) $n=10$; $\bar{x} = 50$; $s^2 = 48.667$; C-I = 99%

a) $t_{0.005, 9} = 3.25$

$$\begin{aligned} \therefore \mu &= \bar{x} \pm t_{0.5\%, 9} \frac{s}{\sqrt{n}} \\ &= 50 \pm 3.25 \sqrt{\frac{48.667}{10}} \\ &= \underline{\underline{57.17, 42.83}} \end{aligned}$$

b) $t_{0.005, 9} = 2.5758$

$$\therefore \mu = 50 \pm 2.5758 \left(\sqrt{\frac{48.667}{10}} \right) = \underline{\underline{55.68, 44.32}}$$

Q8) $\hat{\theta}$ is an estimator of parameter θ .

i) Unbiased Estimator:-

Mean value equals the true parameter value

$$E(\hat{\theta}) = \underline{\underline{\theta}}$$

ii) Bias:-

Bias of an estimator is the difference between this estimator's expected value and the true value of the parameter being estimated. $E(\hat{\theta}) - \theta = 0$

iii) MSE of estimator $\hat{\theta}$.

Measures the average of the squares of the errors - that is, the average squared difference between the estimated values and the actual value.

$$\underline{\underline{\theta = \text{Var} + \text{bias}^2}}$$

iv) $\hat{\theta}$ is more efficient than $\tilde{\theta}$, since MSE of $\hat{\theta}$ is lower than $\tilde{\theta}$.

v) when MSE more towards 0, as the n inverse, estimate would be more

vi) Method of moments

Maximum Likelihood Estimate.

Q10) $n=1000$; $X \sim \text{Poi}(3)$

$$X \sim N(3000, 3000)$$

if $P(X < 3100)$ then $X \sim \text{Poi}(3)$

else X doesn't follow $\text{Poi}(3)$

(i) Type I error = $P(H_0 \text{ is rejected} / H_0 \text{ is true})$

$$= P(X > 3100, \lambda = 3) \therefore X \sim N(3000, 3000)$$

$$= P\left[Z > \frac{3100 - 3000}{\sqrt{3000}}\right]$$

$$= P[Z > 1.826]$$

$$= \underline{3.39\%}$$

(ii) Type 2 error = $P(H_0 \text{ accept} / H_0 \text{ is false})$

(iii) Power = $P(\text{reject } H_0 / H_0 \text{ false})$

$$= 1 - \left[P\left(Z < \frac{3100 - \mu}{\sqrt{\mu}}\right) \right]$$

Q11) $n=12$; $\sum X = 213200$; $\sum X^2 = 4919860000$

$$\bar{X} = 17767 ; S = 10144$$

$$X_1 = 11000 ; X_2 = 48000 ; X_1' = 27500 ; X_2' = 31500$$

$$\sum X = 213200 = A + 11000 + 48000$$

A is sum of the 10 terms.

$$\therefore A = 154200$$

$$\sum X' = A + X_1' + X_2' = 213200$$

$$\therefore \bar{X}' = 17767$$

$$\sum X^2 = 4919860000 = B + 11000^2 + 48000^2$$

$$\therefore B = 2494860000$$

$$\sum X'^2 = B + X_1'^2 + X_2'^2 = 4243360000$$

$$\therefore S = \frac{1}{\sqrt{11}} \sqrt{4243360000 - 12(17767)^2} = \underline{6434.0326}$$

q12) $X_1, X_2, X_3, \dots, X_n \sim U(0, \theta)$; $a=0$; $b=\theta$

(i) $E(Y) = \frac{\theta}{2}$; $V(Y) = \frac{\theta^2}{12}$

$$\therefore L = \prod_{i=1}^n \left(\frac{1}{\theta} \right)$$

$$= \frac{1}{\theta^n}$$

$$\ln(L) = \frac{-n}{\theta}$$

$$\frac{dn}{d\theta} = \frac{-n}{\theta} \text{ when } \theta = \text{infinite}$$

hence CRLB won't be applicable

(ii) $P[\text{Max } X < Y] = \prod_{i=1}^n \left[\int_0^y \frac{dx}{\theta} \right] = \left(\frac{y}{\theta} \right)^n$

$$E[Y^k] = \int y^k \cdot \frac{n}{\theta^n} \cdot y^{n-1} \cdot dy$$

$$= \frac{n}{n+k} \int \frac{n+k}{\theta^n} \cdot y^{n+k-1} \cdot dy$$

$$= \frac{n}{n+k} \left(\frac{\theta^{n+k} - 0}{\theta^n} \right)$$

$$= \frac{n\theta^k}{n+k}$$

(iii) Bias $[\hat{\theta}(c)] = E[\hat{\theta}(c) - \theta]$

$$= E[(y - \theta)]$$

$$= cE(y) - \theta$$

$$= \frac{c \cdot n\theta}{n+1} - \theta$$

$$\therefore \text{Bias}[\hat{\theta}(c)] = \theta \left(\frac{cn}{n+1} - 1 \right)$$

$$(iv) \frac{d}{dc} \text{MSE} [\hat{\theta}(c)] = \frac{d}{dc} \left[\frac{c^2 n \sigma^2}{n+2} - \frac{2n \sigma^2}{n+1} (c \theta) \right]$$

$$C_m = \frac{2n \sigma^2}{n+1} \cdot \frac{n+2}{2n \sigma^2} = \frac{n+2}{n+1}$$

$$\frac{d^2}{dc^2} [\text{MSE} (\hat{\theta}(c))] = \frac{2n \sigma^2}{n-2} > 0 \quad \text{min. value}$$

(vi) $\frac{n+2}{n+1} y \rightarrow y$, two estimators are same

$n \rightarrow \infty$.

