

Probability & Statistics

Assignment-1

1. $X \sim \text{Poi}(\theta)$, where θ is large,
 $\therefore Z = \frac{\hat{\lambda} - \lambda}{\sqrt{\hat{\lambda}/n}} \sim N(0,1)$ approx.

$$\hat{\lambda} = \frac{\sum n_i}{n}, \quad \hat{\lambda} = 200$$

$$P(Z_{2.5\%} < Z < Z_{97.5\%}) = 0.95$$

$$P\left(-1.96 < \frac{\hat{\lambda} - \lambda}{\sqrt{\hat{\lambda}/n}} < 1.96\right) = 0.95$$

$$P\left(-1.96 < \frac{200 - \lambda}{\sqrt{200}} < 1.96\right) = 0.95$$

$$P\left(-1.96\sqrt{200} - 200 < -\lambda < 1.96\sqrt{200} - 200\right) = 0.95$$

$$P\left(200 - 1.96\sqrt{200} < \lambda < 200 + 1.96\sqrt{200}\right) = 0.95$$

$$\therefore 95\% \text{ CI for } \lambda = (200 - 1.96\sqrt{200}, 200 + 1.96\sqrt{200})$$

$$= (172.2814, 227.7186)$$

2.) $X_i \Rightarrow i=1, 2, \dots, n$
 mean = μ , var = σ^2

we know - $\bar{X} = \frac{\sum X_i}{n}$, $s^2 = \frac{1}{n-1} (\sum X_i^2 - n\bar{X}^2)$

(I) $E(\bar{X}) = E\left(\frac{\sum X_i}{n}\right)$

$E(\bar{X}) = \frac{E(\sum X_i)}{n} = \frac{n\mu}{n}$

$\therefore \boxed{E(\bar{X}) = \mu}$

$V(\bar{X}) = V\left(\frac{\sum X_i}{n}\right)$

$V(\bar{X}) = \frac{V(\sum X_i)}{n^2} = \frac{\sum V(X_i)}{n^2}$

$V(\bar{X}) = \frac{n\sigma^2}{n^2}$

$\therefore \boxed{V(\bar{X}) = \frac{\sigma^2}{n}}$

(II) $E(s^2) = E\left[\frac{1}{n-1} (\sum X_i^2 - n\bar{X}^2)\right]$

$E(s^2) = \frac{1}{n-1} E\left[\sum X_i^2 - n\bar{X}^2\right]$

$E(s^2) = \frac{\sum E(X_i^2) - nE(\bar{X}^2)}{n-1}$

we know-

$$V(X) = E(X_i^2) - E(X)^2$$

$$\therefore \sigma^2 = E(X_i^2) - \mu^2$$

$$\therefore E(X_i^2) = \sigma^2 + \mu^2$$

similarly,

$$E(\bar{X}^2) = \mu^2 + \frac{\sigma^2}{n}$$

$$\therefore E(S^2) = \frac{1}{n-1} \left[\sum (\mu^2 + \sigma^2) - n \left(\mu^2 + \frac{\sigma^2}{n} \right) \right]$$

$$E(S^2) = \frac{1}{n-1} \left[n\mu^2 + n\sigma^2 - n\mu^2 - \sigma^2 \right]$$

$$E(S^2) = \frac{1}{n-1} (n-1) \sigma^2$$

$$\boxed{E(S^2) = \sigma^2}$$

Hence proved.

(III)

$$X \sim N(\mu, \sigma^2)$$

$$\therefore \frac{(n-1)S^2}{\sigma^2} = \chi_{n-1}^2$$

$$\text{To find } V(S^2) = \frac{2\sigma^4}{n-1}$$

$$\therefore V\left(\frac{(n-1)S^2}{\sigma^2}\right) = \frac{(n-1)^2}{\sigma^4} \cdot V(S^2)$$

$$\text{But we know, } \frac{(n-1)S^2}{\sigma^2} = \chi_{n-1}^2$$

$$\therefore V(\chi_{n-1}^2) = \frac{(n-1)^2}{\sigma^4} \cdot V(S^2)$$

$$\therefore 2(n-1) = \frac{(n-1)^2}{\sigma^4} \cdot V(S^2)$$

$$\therefore \boxed{V(S^2) = \frac{2\sigma^4}{(n-1)^2}}$$

(IV) $n=10$, $\sigma^2=100$

$$P(50\% \cdot \sigma^2 < S^2 < 50\% \cdot \sigma^2 + \sigma^2)$$

$$= P(0.5\sigma^2 < S^2 < 1.5\sigma^2)$$

$$= P(S^2 < 1.5\sigma^2) - P(S^2 < 0.5\sigma^2)$$

$$= P(S^2 < 150) - P(S^2 < 50)$$

$$= P\left(\frac{9S^2}{\sigma^2} < \frac{150 \times 9}{100}\right) - P\left(\frac{9S^2}{\sigma^2} < \frac{50 \times 9}{100}\right)$$

$$= P\left(\chi^2_9 < 13.5\right) - P\left(\chi^2_9 < 4.5\right)$$

$$= 0.8587 - 0.1245$$

$$\therefore P(50 < S^2 < 150) = 0.7342$$

3.) $X \sim \text{Bin}(4, p)$

(1) $L = \prod_{i=1}^{180} f(u_i)$

$$L = P(X=0)^{86} \times P(X=1)^{75} \times P(X=2)^{16}$$

$$\times P(X=3)^2 \times P(X=4)^1$$

$$L = \left[{}^4C_0 p^0 (1-p)^4 \right]^{86} \times \left[{}^4C_1 p^1 (1-p)^3 \right]^{75} \times \left[{}^4C_2 p^2 (1-p)^2 \right]^{16}$$

$$\times \left[{}^4C_3 p^3 (1-p) \right]^2 \times \left[{}^4C_4 p^4 (1-p)^0 \right]^1$$

$$L = C \times p^{117} \times (1-p)^{603}$$

where C is constant.

$$\log L = \log(C) + 117 \log(p) + 603 \log(1-p)$$

differentiating wrt p -

$$\frac{dL}{dp} = 0 + \frac{117}{p} - \frac{603}{(1-p)}$$

$$\therefore \frac{117}{p} - \frac{603}{(1-p)} = 0$$

$$117 - 117p = 603p$$

$$\therefore p = \frac{117}{720}$$

$$\therefore \boxed{p = 0.1625}$$

(11) ~~33~~

H_0 = no. of claims conform to binomial

H_1 = no. of claims not conform binomial.

	0	1	2	3	4
O_i	86	75	16	2	1
E_i	88.542	68.724	19.998	2.574	0.108

~~$$P(X=0) = \frac{1}{6}$$~~

$$P(X=n) = {}^n C_n p^n (1-p)^{n-n}$$

$$\therefore P(X=0) = 0.4919$$

$$P(X=1) = 0.3818$$

$$P(X=2) = 0.1111$$

$$P(X=3) = 0.0143$$

$$P(X=4) = 0.0006$$

O_i	86	75	19
E_i	88.542	68.724	22.68

$$T.S. = \frac{\sum (O_i - E_i)^2}{E_i} \sim \chi^2_1$$

$$T.S. = 1.243223$$

$\therefore$ we can accept that it follows binomial model.

$$4) f(n) = \frac{c\beta^3}{(n+\beta)^4}, \quad n > 0, \beta > 0$$

(I) Given P.P.F looks like pareto distribution with PDF-

$$\text{PDF} = \frac{\alpha \lambda^\alpha}{(\lambda + n)^{\alpha+1}}$$

$$\therefore \beta = \lambda \text{ here}$$

$$\therefore c = \alpha \text{ here}$$

$$\therefore \cancel{c=2} \quad \boxed{c=3}$$

$$E(x) = \frac{\lambda}{\alpha-1} = \frac{\beta}{2}$$

$$V(x) = \frac{\alpha \lambda^2}{(\alpha-1)^2(\alpha-2)} = \frac{3\beta^2}{4}$$

$$(II) E(x) = \frac{\beta}{2}$$

$$\therefore \bar{x} = \frac{\beta}{2}$$

$$\therefore \hat{\beta} = 2\bar{x}$$

Hence proved.

Bias:

$$\text{Bias} = E(\hat{\beta}) - \beta$$

$$\text{Bias} = E(2\bar{x}) - \beta = 2 E(\bar{x}) - \beta$$

$$\text{Bias} = \frac{2}{n} E(\sum x_i) - \beta = \frac{2}{n} \times \frac{n\beta}{2} - \beta$$

$$\boxed{\text{Bias} = 0}$$

Consistency -

$$\text{MSE} = V(\hat{\beta}) + \text{Bias}^2$$

$$\text{MSE} = V(\hat{\beta})$$

$$\text{MSE} = V(2\bar{x})$$

$$\text{MSE} = 4 V(\bar{x})$$

$$\text{MSE} = \frac{4}{n^2} \sum V(x_i) = \frac{4}{n^2} \frac{3\beta^2 n}{4}$$

$$\text{MSE} = \frac{3\beta^2}{n}$$

$\therefore$ as n increases, MSE closes to 0

$\therefore \hat{\beta}$ is consistent.

(III)

$$\hat{\beta} = b\bar{x}$$

$$E(\hat{\beta}) = E(b\bar{x})$$

$$= b E\left(\frac{\sum x_i}{n}\right) = \frac{b}{n} \sum E(x_i)$$

$$E(\hat{\beta}) = \frac{b\beta}{2}$$

$$\begin{aligned} V(\hat{\beta}) &= V(b\bar{x}) \\ &= \frac{b^2}{n^2} \sum V(x_i) \end{aligned}$$

$$= \frac{3b^2\beta^2}{4n}$$

$$\begin{aligned} \text{MSE} &= V(\beta) + \text{Bias}^2 \\ &= \frac{3b^2\beta^2}{4n} + \left(\frac{b\beta - 2\beta}{2} \right)^2 \\ &= \frac{3b^2\beta^2}{4n} + \frac{b^2\beta^2 + 4\beta^2 - 4\beta^2 b}{4} \end{aligned}$$

$$\text{MSE} = \frac{\beta^2}{4n} (3b^2 + nb^2 + 4n - 4nb)$$

diff. wrt b

$$\frac{dM}{db} = \frac{\beta^2}{4n} (6b + 2nb - 4n)$$

$$\frac{\beta^2}{4n} (6b + 2nb - 4n) = 0$$

$$\therefore 6b + 2nb - 4n = 0$$

$$b(3+n) = 2n$$

$$\therefore b = \frac{2n}{3+n} = \frac{2}{\left(1 + \frac{3}{n}\right)}$$

$$\frac{d^2M}{db^2} = \frac{\beta^2}{4n} (6+2n) > 0$$

$$\therefore b = \frac{2}{\left(1 + \frac{3}{n}\right)} \text{ minimizes MSE}$$

5.)

(i) $X \sim U(a, b)$

$$E(X) = \frac{a+b}{2} = \bar{X}$$

$$\therefore b = 2\bar{X} - a \quad \text{--- (1)}$$

$$V(X) = \frac{(b-a)^2}{12} = S^2$$

$$\therefore 12S^2 = (b-a)^2$$

$$\therefore 12\sqrt{3}S = b-a$$

$$12\sqrt{3}S = 2\bar{X} - 2a$$

$$\therefore \sqrt{3}S = \bar{X} - a$$

$$\therefore \boxed{\hat{a} = \bar{X} - \sqrt{3}S}$$

(ii) $\hat{b} = 2\bar{X} - a$

$$\therefore \boxed{\hat{b} = \bar{X} + \sqrt{3}S}$$

(iii) sample observations - $\{10, 20, 30, 40, 50\}$

$$\bar{X} = 30, S =$$

$$6. > \quad H_0 = p = 0.1$$

$$H_1 = p > 0.1$$

$$\begin{aligned}
 (i) \quad P(\text{type I error}) &= P(\text{rejecting } H_0 / H_0 \text{ is true}) \\
 &= P(X \geq 3 | p=0.1) + P(X=0 | p=0.1) \cdot P(Y \geq 5 | p=0.1) \\
 &\quad + P(X=1 | p=0.1) \cdot P(Y \geq 4 | p=0.1) \\
 &\quad + P(X=2 | p=0.1) \cdot P(Y \geq 3 | p=0.1) \\
 &= (1 - 0.8891) + (0.2824 \times 0.0043) \\
 &\quad + (0.3766 \times 0.2566) + (0.2301 \times 0.1109) \\
 &= 0.147268 \approx 15\%.
 \end{aligned}$$

$$\begin{aligned}
 (ii) \quad &\text{P(rejecting when } p=0.3) = P(X \geq 3 | p=0.3) + \\
 &\quad P(X=0 | p=0.3) P(Y \geq 5 | p=0.3) \\
 &\quad + P(X=1 | p=0.3) P(Y \geq 4 | p=0.3) \\
 &\quad + P(X=2 | p=0.3) P(Y \geq 3 | p=0.3) \\
 &= 0.912532
 \end{aligned}$$

$$\begin{aligned}
 (iii) \quad P(\text{type II error}) &= 1 - 0.912532 \\
 &= 0.87468
 \end{aligned}$$

$$(iv) \frac{X/n - p}{\sqrt{p(1-p)}} \sim N(0, 1)$$

$$n=120, X=40$$

$$\hat{p} = \frac{40}{120} = \frac{1}{3} = 0.3333$$

$$\therefore Z_{10\%} = 1.2816$$

$$\hat{p} - 1.2816 \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = 0.3333 - 1.2816 \sqrt{\frac{0.3333(0.6667)}{120}}$$

$$= 0.2782$$

$$(v) p > 0.2782 \text{ at } 10\% \text{ L.O.S.}$$

$$H_0 \Rightarrow p = 0.2782, H_1 \Rightarrow p > 0.2782$$

$$\frac{X - np_0}{\sqrt{np_0q_0}} = \frac{39.5 - 120(0.2782)}{\sqrt{120 \times 0.2782 \times 0.7218}}$$

$$= 1.2511$$

$\therefore$ we can not reject H_0

(vii)		I_1	I_2
n		12	15
$\bar{X}$		65	70
S_1		54	70

$$S_p^2 = \frac{1}{25} (11(2916) + 14(4900)) = 4027.04$$

$$\frac{65 - 70}{\sqrt{\frac{4027.04}{25}}} = -0.2034$$

$\therefore \pm 2.060$ at 5% level,

7.) (i)	1	2
	Public	Private
$\bar{X}$	30	35.0555
S^2	64.3125	67.4027

(ii) if populations have equal variance, t-test can be used.

(iii) $H_0 \Rightarrow \mu_1 = \mu_2$
 $H_1 \Rightarrow \mu_2 > \mu_1$

$$TS = \frac{(\bar{X}_2 - \bar{X}_1) - (\mu_2 - \mu_1)}{SP \sqrt{\frac{n_1 + n_2}{n_1 n_2}}}$$

$$TS = \frac{35.0555 - 30}{8.1152 \sqrt{\frac{2}{9}}}$$

$$TS = 11.32151$$

tabulated value $t_{16, 5\%} = 1.746$

$\therefore TS < \text{tabulated value}$

$\therefore H_0$ is not rejected.

8.) (i) unbiased estimator : it is an estimator whose expected value is equal to parameter. It is possible when bias is equal to 0

(ii) bias : it is the positive or negative difference in the estimated and actual value.

(iii) $MSE(\hat{\theta}) = \text{var}(\theta) + \text{bias}^2$

(iv) $\tilde{\theta}$ is more efficient as MSE value is lower compared to $\hat{\theta}$.

(v) ~~Estimators~~ Estimators will be called consistent when their value is close to 0. In this case $\tilde{\theta}$ is ^{more} consistent than $\hat{\theta}$.

(vi) Maximum likelihood estimate and method of moments.

9.) (i) $t_k \Rightarrow \frac{N(0,1)}{\sqrt{\chi_k^2/k}}$ where k is degree of freedom.

(ii) $k > 2$

$$E(t_k) = 0, \quad V(t_k) = \frac{k}{k-2}$$

(iii) [a] $\frac{\bar{x} - \mu}{S/\sqrt{n}} \sim t_{n-1}, \quad n = 10$

$$P\left(t_{9,0.5\%} < \frac{\bar{x} - \mu}{S/\sqrt{n}} < t_{9,99.5\%}\right) = 0.99$$

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$$P\left(\bar{X} - t_{9, 0.5\%} \cdot \frac{S}{\sqrt{n}} < \mu < \bar{X} + t_{9, 0.5\%} \cdot \frac{S}{\sqrt{n}}\right) \geq 0.99$$

$$\therefore t_{9, 0.5\%} = 3.250$$

$\therefore$ 99% confidence interval -

$$\mu = \left(50 - \frac{3.25 \cdot 48.667}{\sqrt{10}}, 50 + \frac{3.25 \cdot 48.667}{\sqrt{10}} \right)$$

$$\mu = (42.83, 57.17)$$

$$[b] \quad t_{n-1} \sim N\left(0, \sqrt{\frac{n-1}{n-3}}\right)$$

$$t_{n-1} \sim N\left(0, \sqrt{\frac{9}{7}}\right)$$

$$\therefore \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim N\left(0, \sqrt{\frac{9}{7}}\right)$$

$$\therefore \frac{\bar{X} - \mu}{S/\sqrt{7 \times 9}} \sim N(0, 1)$$

critical value : 2.58

$$\begin{aligned} \text{confidence interval} &= \left(50 - 2.58 \sqrt{\frac{48.667 \times 9}{10 \times 7}}, 50 + 2.58 \sqrt{\frac{48.667 \times 9}{10 \times 7}} \right) \\ &= (43.5463, 56.4535) \end{aligned}$$

10.) $n = 1000$
 $\lambda = 3$
 $X \sim \text{Poi}(n\lambda) ; \text{Poi}(3000)$

(i) $P(\text{type I error}) = P(\text{reject } H_0 \mid H_0 \text{ is true})$
 $= P(u > 3100 \mid X \sim \text{Poi}(3000))$

$$\therefore P(u > 3100) = P\left(\frac{Z > \frac{3100 - 3000}{\sqrt{3000}}}\right)$$

$$= 1 - P(Z < 1.825)$$

$$P(u > 3100) = 0.0339$$

(ii) Type II error is when hypothesis is false.

(iii) Power of test $= P(\text{rejecting } H_0 \mid H_0 \text{ is false})$
 $= P(u > 3100 \mid X \sim \text{Poi}(n\lambda) \sim N(n\lambda, n\lambda))$
 $= 1 - P\left(Z < \frac{3100 - n\lambda}{\sqrt{n\lambda}}\right)$

(iv) ~~SE~~
 $P\left(-2.5758 < \frac{2.9 - u}{\sqrt{2.9/1000}} < 2.5758\right) = 0.99$

$\therefore$ confidence interval for

$$u = (2.7613, 3.0387)$$

11.)

$$n=12, \sum X = 213200, S=10,144.$$

$$\sum X^2 = 4919860000, \bar{X} = 17,767$$

$$\begin{aligned} \sum X_1 &= \sum X - 11000 - 48000 + 27500 + 31500 \\ &= 213200 \end{aligned}$$

$$\text{new } \bar{X} = \frac{213200}{12}$$

$$\boxed{\text{new } \bar{X} = 17,767}$$

$$\begin{aligned} \sum X_1^2 &= 4919860000 - (11000)^2 - (48000)^2 \\ &\quad + (27500)^2 + (31500)^2 \\ \sum X_1^2 &= 4243360000 \end{aligned}$$

$$\begin{aligned} \text{new } S^2 &= \frac{1}{11} (4243360000 - 12(17767)^2) \\ &= 41396775.64 \end{aligned}$$

$$\therefore \boxed{\text{new } S = 6434.032611}$$