

PROBABILITY & STATISTICS - 2ASSIGNMENT - 2

1. (a) Let regression line y on x be

$$\hat{y} = \hat{\alpha} + \hat{\beta}x$$

$$\hat{\beta} = \frac{S_{xy}}{S_{xx}} = \frac{\sum xy - n\bar{x}\bar{y}}{\sum x^2 - n\bar{x}^2}$$

$$\sum xy = 182560$$

$$\bar{x} = \frac{\sum x_i}{n} = 85.0$$

$$\bar{y} = \frac{\sum y_i}{n} = ~~137~~ 173.7$$

$$\sum x^2 = 92500$$

$$\Rightarrow \hat{\beta} = \frac{34915}{20250} = \underline{\underline{1.72421}}$$

$$\therefore \hat{y} = 27.143 + 1.7242x$$

(b) Since the slope (β) is positive, y and x have a positive relation.

Q2. (i) $\hat{\beta} = \frac{S_{xy}}{S_{xx}}$ and $\hat{\alpha}_i = \bar{y} + \hat{\beta} \bar{x}_i$

$$S_{xx} = \sum x_i^2 - \frac{(\sum x_i)^2}{n} = 12666.58 - \frac{(385.2)^2}{12}$$

$$= 301.66$$

$$S_{yy} = \sum y_i^2 - \frac{(\sum y_i)^2}{n}$$

$$= 119026.9 - \frac{(1162.5)^2}{12} = 6409.7125$$

$$S_{xy} = \sum x_i y_i - \frac{(\sum x_i)(\sum y_i)}{n}$$

$$= 38191.41 - \frac{(385.2)(1162.5)}{12}$$

$$= 875.16$$

$$\hat{\alpha}_i = 3.748181 \quad \hat{\beta} = 2.901146$$

$$\hat{y}_i = 3.748181 + 2.901146 x_i$$

(ii) $PQ: \hat{\beta} - \beta \sim t_{n-2}$

$$se(\hat{\beta}) = \sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}$$

$$\hat{\sigma}^2 = \frac{1}{10} \left[6409.7125 - \frac{(875.16)^2}{301.66} \right] = 387.0744$$

$$se(\hat{\beta}) = 1.132$$

$$P[-2.228 < \frac{2.901 - \beta}{1.132} < 2.228] = 0.95$$

$$P[-2.228(1.132) - 2.901 < -\beta < 2.228(1.132) - 2.901] = 0.95$$

$$95\% \text{ CI for } \beta = \left\{ 2.901 - 2.228(1.132), 2.901 + 2.228(1.132) \right\}$$

$$= \{ 0.378904, 5.423096 \}$$

$$(iii) \quad PQ = \frac{\hat{\mu}_i - \mu_i}{Se(\hat{\mu}_i)} \sim t_{n-2}$$

$$\hat{\mu}_i = 3.748181 + 2.901146(10) \\ = 119.79$$

$$Se(\mu_i) = \sqrt{\left[\frac{1}{12} + \frac{(9.9)^2}{301.66} \right] 387.07} = 10.5988$$

$$P[-2.228 < \frac{119.79 - \mu_i}{10.5988} < 2.228] = 0.95$$

$$95\% \text{ CI for } \mu_i = \left\{ 119.79 - 2.228(10.5988), 119.79 + 2.228(10.5988) \right\} \\ = \{ 96.17, 143.40 \}$$

Q3. Comments

(i) The plot suggest:

The variation in city C is least as compared to which has the maximum variance

→ The time difference in commutation in peak hours and non-peak hours is maximum in city A and least in city D.

(ii) Assumptions:

- The population ~~means~~ has a common variance
- The observations are independent
- The population must be normal

(iii) H_0 : The mean of difference is same for each city

H_1 : The mean of difference is not the same

$$SS_{TOT} = S_{yy} = \sum y_i^2 - \frac{(\sum y_i)^2}{n}$$

$$= 1495 - \frac{103^2}{28} = 1116.11$$

$$SS_{REG} = \frac{1}{7} (64^2 + 44^2 + 8^2 + (-13)^2) - \frac{103^2}{28} = 516.11$$

$$SS_{RES} = SS_{TOT} - SS_{REG} = 600$$

Anova Table

Sr of variation	df	Sum of sq.	mean sum of sq.
Regression	3	516.11	172.04
Residual	24	600	25
Total	27	1116.11	197.04

$$\text{Under } H_0, F = \frac{MS_{REG}}{MS_{RES}} = \frac{172.04}{25} = 6.88$$

$$\text{Test Statistic} = 6.88$$

$$F_{3, 24, 5\%} = 3.009 \quad \{ \text{Right tailed test} \}$$

Hence we have sufficient evidence to reject H_0 at 5% level. Thus, there are underlying difference b/w the cities.

IV Analysis of mean differences

Since, $\bar{y}_1 = 9.14$

$$\bar{y}_2 = 6.29$$

$$\bar{y}_3 = 1.14$$

$$\bar{y}_4 = -1.86$$

$$\bar{y}_1 > \bar{y}_2 > \bar{y}_3 > \bar{y}_4$$

$$\hat{\sigma}^2 = \frac{SS_{RES}}{n-2} = 25$$

The least significance difference between mean is

$$t_{24}, 0.025 \times \hat{\sigma}^2 \sqrt{\frac{1}{7} + \frac{1}{7}} = 2.064 \times \sqrt{25} \times \sqrt{\frac{1}{7} + \frac{1}{7}} = 5.52$$

$$\bar{y}_1 - \bar{y}_2 = 2.85$$

$$\bar{y}_2 - \bar{y}_3 = 5.15$$

$$\bar{y}_3 - \bar{y}_4 = 3$$

All are less than 5.52

$$\bar{y}_1 > \bar{y}_2 > \bar{y}_3 > \bar{y}_4$$

Q4. (a) Mathematical model :

$$y_{ij} = \mu + \tau_i + e_{ij} \text{ where}$$

y_{ij} = j^{th} observation on the i^{th} treatment

μ = common effect for the whole experiment

τ_i = i^{th} treatment effect

e_{ij} = random error in the j^{th} observation on the i^{th} treatment

assumptions

→ errors are assumed to be normally & independently distributed with mean 0 & constant variance σ^2

→ μ is fixed parameter

(b) H_0 : mean claims are equal

H_1 : mean claims are not equal

$$m_1 = 7, m_2 = 8, m_3 = 6, m_4 = 7, m_5 = 5$$

$$m_{tot} = 33$$

$$\sum_{i=1}^5 \sum_{j=1}^{m_i} y_{ij} = 354 + 386 + 87 + 648 + 384 = 1856$$

$$\sum_{i=1}^5 \sum_{j=1}^{m_i} y_{ij}^2 = 27184 + 29430 + 2123 + 84669 + 30910 = 174316$$

$$SS_{TOT} = 174316 - \left[33 \times \left(\frac{1856}{33} \right)^2 \right]$$

$$= 69930.06$$

$$SS_{REG} = \left[\frac{354^2}{7} + \frac{386^2}{8} + \frac{87^2}{6} + \frac{648^2}{7} + \frac{384^2}{5} \right] - \left[33 \times \left(\frac{1856}{33} \right)^2 \right]$$

$$= 22325.69$$

$$SS_{RES} = SS_{TOT} - SS_{REG} \\ = 47604.37$$

Src of var.	DOF	Sum of sq	MSS
Regression	4	22325.69	5581.4225
Residual	28	47604.37	1700.156
Total	32	69930.06	

$$F_s = \frac{MSS_{REG}}{MSS_{RES}} \sim f_{4,28}$$

$$F_{cal} = \frac{5581.4225}{1700.156} = 3.2829$$

$$F_{tab} = F_{4,28} = 2.714$$

Since $F_{cal} > F_{tab}$ we reject H_0
Mean claims are not equal for the 5 companies

(c) H_0 : No association between salaries & no. of papers cleared

H_1 : There is an interdependence between salary and papers cleared

O_i salaries papers	3-5	5-8	8-10	10-12	Total
0-3	45	20	6	5	76
4-6	7	20	9	6	42
7-9	5	8	15	12	40
Total	57	48	30	23	158

E_i

	3-5	5-8	8-10	10-12
0-3	27.41772	23.08861	14.43038	11.06329
4-6	15.15189	12.78949	9.97468	6.11392
7-9	14.43038	12.15189	7.59494	5.82278

$$\chi^2 = \sum \frac{(O_i - E_i)^2}{E_i} \sim \chi^2_{(3-1)(2-1)}$$

$$\chi^2_{cal} = 49.91146$$

$$\chi^2_{tab} = \chi^2_{6.5\%} = 12.59$$

Since $\chi^2_{cal} > \chi^2_{tab}$, we reject H_0

Salaries and no. of papers cleared are associated.

Q5. (i) Assumptions

- Populations are normal
- Population have common variance
- Observation are independent

The sample variance are very different from each other. Hence, the common variance assumption won't hold.

(ii) Missing values are:

$$\text{Mean} \rightarrow \frac{\sqrt{29} + \sqrt{3} + \sqrt{21}}{3} = 4.5244$$

$$\text{Var} \rightarrow \frac{\sum (x_i - \bar{x})^2}{n-1} = 0.1213$$

(iii) log transformation is the best suited as the common variance assumption holds true in this case

$$(iv) (a) \quad y_{ij} = \mu + \tau_i + \varepsilon_{ij} \quad ; \quad i = 1, 2, 3, \dots \\ j = 1, 2, 3, \dots$$

$$H_0 : \tau_i = 0$$

$$H_1 : \tau_i \neq 0$$

$$(b) \quad y_i^2 = \sum y_{ij}^2 = \frac{(3 \times \text{mean}_i)^2}{3} = 973.373$$

$$SS_{\text{reg}} = \sum_{i=1}^4 \frac{y_i^2}{3} - \frac{y^2}{12} = 12.153$$

$$SS_{\text{reg}} = \sum_{i=1}^4 \left(\frac{2}{j=1} (y_{ij} - \bar{y}_i)^2 \right)$$

$$= \sum_{i=1}^4 (2 \times \text{variance})$$

$$= 2 \times 0.618$$

$$= 1.236$$

Src of vari	Dof	sum of sq	MSS
Regression	3	21.153	7.051
Residual	8	1.236	0.155
Total	11	22.386	

$$T_b: \frac{MSS_{reg}}{MSS_{res}} \sim F_{3,8}$$

$$F_{cal} = 45.638$$

$$F_{tab} = F_{3,8, 5\%} = 4.066$$

Since $F_{cal} > F_{tab}$, we reject H_0

$$\begin{aligned}
 6. \text{ (i) correlation coeff } (r) &= \frac{S_{xy}}{\sqrt{S_{xx} S_{yy}}} \\
 &= \frac{\sum xy - n\bar{X}\bar{Y}}{\sqrt{(\sum x^2 - n\bar{X}^2)(\sum y^2 - n\bar{Y}^2)}} \\
 &= \frac{661.2}{\sqrt{34.9 \times 8923.6}} \\
 &= \underline{\underline{0.9447}}
 \end{aligned}$$

$$(ii) \hat{\beta} = \frac{S_{xy}}{S_{xx}} = \frac{661.2}{54.9} = 12.04371$$

$$\begin{aligned}
 \hat{\alpha} &= \bar{Y} - \hat{\beta}\bar{X} = 56.2 - 12.04371 \times 3.9 \\
 &= 9.2295
 \end{aligned}$$

$$(iii) SSTOT = SS_Y = 8923.6$$

$$SS_{REG} = \frac{S^2_{xy}}{S_{xx}} = \frac{661.2^2}{54.9} = 7963.30492$$

$$SS_{RES} = SSTOT - SS_{REG} = 960.29508$$

$$\text{in } R^2 = \frac{S^2_{xy}}{S_{xx} \cdot S_{yy}} = 89.24\%$$

$$\text{Q7. (i) } \sum x = 174.13 \quad \sum y = 199.84$$

$$\sum x^2 = 7545.9 \quad \sum y^2 = 4747.45$$

$$\sum (x - \bar{x})(y - \bar{y}) = 2176.84$$

$$n = 1$$

$$S_{xx} = 7545.9 - \frac{174.13^2}{11}$$

$$= 4789.4221$$

$$S_{yy} = 4747.45 - \frac{199.84^2}{11}$$

$$= 1189.20769$$

$$S_{xy} = 2176.84$$

$$\hat{\beta} = \frac{S_{xy}}{S_{xx}} = 0.45451$$

$$\hat{\alpha} = \bar{y} - \hat{\beta} \bar{x} = 10.78056$$

$$\hat{y} = 10.78056 + 0.45451x$$

$$\text{(ii) } H_0: \beta = 0 \quad H_1: \beta \neq 0$$

$$T.S \rightarrow \frac{\hat{\beta} - \beta}{S_e(\hat{\beta})} \sim t_{n-2}$$

$$\hat{\sigma}^2 = \frac{1}{n-2} \left(S_{yy} - \frac{S_{xy}^2}{S_{xx}} \right) = 22.20136$$

$$S_e(\hat{\beta}) = \sqrt{\frac{\hat{\sigma}^2}{S_{xx}}}$$

$$T.S = \frac{0.45451 - 0}{\sqrt{\frac{22.20136}{4789.4221}}} \sim t_9$$

$$6.6755 = t_9$$

Assumptions

- ① Popⁿ have a common variance
- ② Popⁿ follow the normal distⁿ
- ③ The obsⁿ are independent

$$(iii) \quad r = \frac{S_{xy}}{\sqrt{S_{xx}S_{yy}}} = 0.91213$$

$$(iv) \quad X_0 = 25$$

Ind. response

$$Y_0 \sim N [\hat{\alpha} + \hat{\beta} X_0, \hat{\sigma}^2 \left[1 + \frac{1}{n} + \frac{(X_0 - \bar{x})^2}{S_{xx}} \right]]$$

$$TS : \frac{Y_0 - (\hat{\alpha} + \hat{\beta} X_0)}{\hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(X_0 - \bar{x})^2}{S_{xx}}}} \sim t_{n-2}$$

$$\begin{aligned} 95\% \text{ CI} &= Y_0 \pm t_{n-2, 0.025} \sqrt{\hat{\sigma}^2 \left(1 + \frac{1}{n} + \frac{(X_0 - \bar{x})^2}{S_{xx}} \right)} \\ &= (\hat{\alpha} + \hat{\beta} X_0) \pm t_{9, 0.025} \sqrt{\hat{\sigma}^2 \left(1 + \frac{1}{n} + \frac{(X_0 - \bar{x})^2}{S_{xx}} \right)} \\ &= (10.9319, 33.3745) \end{aligned}$$

Mean response

$$TS : \frac{\mu_0 - (\hat{\alpha} + \hat{\beta} X_0)}{\hat{\sigma} \sqrt{\frac{1}{n} + \frac{(X_0 - \bar{x})^2}{S_{xx}}}} \sim t_{n-2}$$

$$\begin{aligned} 95\% \text{ CI} &= 2 + \hat{\beta} X_0 \pm t_{9, 0.025} \sqrt{\hat{\sigma}^2 \left(\frac{1}{n} + \frac{(X_0 - \bar{x})^2}{S_{xx}} \right)} \\ &= (18.643, 25.663) \end{aligned}$$

- (v) The residual plot appears to follow a pattern (first increasing and then decreasing), which implies that the model is a good fit for the data.

8. (i) Factor analysis provides a method for reducing the dimensionality of the data set i.e. it seeks to identify the key components necessary to model and understand data.

→ the original data is rewritten in terms of few variables that are unrelated; only some are needed to explain most of the variability observed in the data.

The new variables are called principle components.

(ii) variance-cov matrix of PC1, PC2, PC3, PC4, PC5 is given by:

$$\begin{bmatrix} 0.456 & 0 & 0 & 0 & 0 \\ 0 & 0.137 & 0 & 0 & 0 \\ 0 & 0 & 0.080 & 0 & 0 \\ 0 & 0 & 0 & 0.0165 & 0 \\ 0 & 0 & 0 & 0 & 0.012 \end{bmatrix}$$

% of total variance explained by each PC

$$PC_1 = \frac{0.456}{0.7015} \times 100 = 65\%$$

$$PC_2 = \frac{0.137}{0.7015} \times 100 = 19.53\%$$

$$PC_3 = \frac{0.080}{0.7015} \times 100 = 11.40\%$$

$$PC_4 = \frac{0.0165 \times 100}{0.7015} = 2.35\%$$

$$PC_5 = \frac{0.012 \times 100}{0.7015} = 1.71\%$$

ii) We can conclude that the dimensionality could be reduced 3. using the first 3 PC's, as 96% of the variability has been explained by these 3 alone.

9 (i) Based on the plot, there is a negative linear correlation between marks and hours spent/day on social media.

$$\text{cii) Correlation Coeff (r)} = \frac{S_{xy}}{\sqrt{S_{xx} S_{yy}}}$$

$$= \frac{\sum xy - n\bar{x}\bar{y}}{\sqrt{(\sum x^2 - n\bar{x}^2)(\sum y^2 - n\bar{y}^2)}}$$

$$= \frac{-296}{\sqrt{75 \times 2482.4}}$$

$$= \underline{\underline{-0.6860}}$$

$$\text{iii) } H_0: \rho = 0$$

$$H_1: \rho < 0$$

Using Fisher's Transformation,

$$w = \frac{1}{2} \ln \left(\frac{1+r}{1-r} \right) \ln \left(\frac{1+\rho}{1-\rho} \right), \frac{1}{n-3}$$

$$Z_{cal} = \frac{-0.84036 - 0}{\sqrt{\frac{1}{7}}} = -2.2234$$

$$Z_{tab} = -1.98$$

Since $Z_{cal} < Z_{tab}$; we reject H_0 . Thus there is negative correlation between the variables.

(iv) let the linear regression model :

$$\hat{y} = \hat{\alpha} + \hat{\beta} x$$

$$\hat{\beta} = \frac{S_{xy}}{S_{xx}} = \frac{-296}{75} = -3.9467$$

$$\hat{\alpha} = \bar{y} - \hat{\beta} \bar{x}$$

$$= 64.4 - (-3.9467)(4.5)$$

$$= \underline{82.16}$$

$$\therefore \hat{y} = \underline{82.16 - 3.9467x}$$

$$(v) \text{ Coeff of determination } (R^2) = \frac{S^2_{xy}}{S_{xx} S_{yy}} = \frac{(-246)^2}{75 \times 24824}$$

$$= \underline{0.4706}$$

$$\underline{r^2 = R^2}$$

$$vi) \text{ expected change in marks} = \underline{\underline{\beta \Delta x}} \\ = \underline{\underline{-3.9467}}$$

10. H_0 : Industry has no impact on resignation rate
 H_1 : There is an impact

$$\text{Mean} = \frac{27 + 36 + 30}{3} = 31$$

$$SS_{\text{Reg}} = 20 \left[(27 - 31)^2 + (36 - 31)^2 + (30 - 31)^2 \right] = 840$$

$$SS_{\text{Reg}} = 19(5^2 + 10^2 + 8^2) = 3591$$

ANOVA TABLE

Source of variation	DOF	Sum of Sq	MSS
Residual	57	3591	63
Regression	2	840	420

$$TS : \frac{MSS_{\text{Reg}}}{MSS_{\text{Res}}} \text{ in } F_{2, 57}$$

$$F_{\text{cal}} = \frac{420}{63} = 6.67$$

$$F_{\text{tab}} \approx F_{2, 60.5} = 3.150$$

→ Since $F_{\text{cal}} > F_{\text{tab}}$, we reject H_0 . There is an impact of industry on the resignation rate.