

Assignment-1

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Subject: Probability & Statistics-2

Section: FY-B.

Solⁿ ① $X \sim \text{Pois}(\theta)$

$$\frac{X-\theta}{\sqrt{\theta}} \sim N(0,1)$$

$$P[X_1 < 0 < X_2] = 0.95$$

$$X \sim \text{Pois}(200)$$

$$\begin{aligned} 95\% \text{ CI} &= \mu \pm 1.96\sqrt{\sigma} \\ &= 200 \pm 1.96\sqrt{200} \end{aligned}$$

Hence the interval is (173.388, 227.7186)

Solⁿ ② $\bar{X} = \sum_{i=1}^n X_i$: sample mean

$$S^2 = \frac{1}{(n-1)} \sum_{i=1}^n (X_i - \bar{X})^2$$
 : sample variance

$$\text{i) } E(\bar{X}) = E\left[\frac{\sum X_i}{n}\right] = \frac{1}{n} (\sum X_i)E = \frac{1}{n} [\sum E(X_i)] = \frac{n\mu}{n} = \mu$$

$$V(\bar{X}) = V\left[\frac{\sum X_i}{n}\right] = \frac{1}{n^2} V[\sum X_i] = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

$$\begin{aligned} \text{(ii) } E(S^2) &= E\left[\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2\right] \\ &= \frac{1}{n-1} E[\sum X_i^2 - n\bar{X}^2] \end{aligned}$$

$$= \frac{1}{n-1} [E(\sum X_i^2) - E(n\bar{X}^2)] = \frac{1}{n-1} [E[\sigma^2 + \mu^2] - n\left[\frac{\sigma^2}{n} + \mu^2\right]]$$

$$\frac{1}{n-1} [n\sigma^2 + n\bar{x}^2 - \sigma^2 + \sigma^2]$$

$$= \frac{1}{n-1} [(n-1)\sigma^2] = \boxed{\sigma^2}$$

Now here $\frac{(n-1)s^2}{\sigma^2} \sim \chi^2_{n-1}$

$$\text{Var} \left[\frac{(n-1)s^2}{\sigma^2} \right] = \chi^2_{n-1}$$

$$\frac{(n-1)^2 \cdot \text{Var}[s^2]}{\sigma^4} = 2(n-1)$$

$$\boxed{\text{Var}[s^2] = \frac{2\sigma^4}{(n-1)}}$$

(iv) $n=10$

$$\sigma^2 = 100.$$

$$\frac{(n-1)s^2}{\sigma^2} = 0.5$$

$$\frac{9 \times s^2}{100} = 0.5$$

$$s^2 = \frac{50000}{9}$$

$$\boxed{s^2 = 5.56:1}$$

③

No. of Claim	0	1	2	3	4
No. of Policies	86	75	16	2	1

(i) MLE

$$X \sim \text{Bin}(4, p)$$

$$L = P(X=0)^{86} \times P(X=1)^{75} \times P(X=2)^{16} \times P(X=3)^2 \times P(X=4)^1$$

$$= \left[{}^4C_0 \cdot p^0 \cdot (1-p)^4 \right]^{86} \times \left[{}^4C_1 \cdot p^1 \cdot (1-p)^3 \right]^{75} \times \left[{}^4C_2 \cdot p^2 \cdot (1-p)^2 \right]^{16} \times \left[{}^4C_3 \cdot p^3 \cdot (1-p)^1 \right]^2 \times \left[{}^4C_4 \cdot p^4 \cdot (1-p)^0 \right]^1$$

$$= \left[\text{Cons.} \times (1-p)^4 \right]^{86} \times \left[\text{Cons.} \times p \cdot (1-p)^3 \right]^{75} \times \left[\text{Cons.} p^2 (1-p)^2 \right]^{16} \times \left[\text{Cons.} p^3 (1-p) \right]^2 \times \left[{}^4C_4 p^4 \right]^1$$

$$\ln(L) = 344 \ln(1-p) + 75 \ln(p) + 256 \ln(1-p) + 32 \ln p + 32 \ln(1-p) + 6 \ln p + 2 \ln(1-p) + 4 \ln p$$

$$\ln(L) = 633 \ln(1-p) + 117 \ln p$$

$$\frac{d[\ln(L)]}{dp} = \frac{117}{p} - \frac{633}{1-p}$$

$$\text{let } \frac{d[\ln(L)]}{dp} = 0$$

$$\frac{117}{p} = \frac{633}{1-p}$$

$$117 - 117p = 633p$$

$$117 = 750p$$

$$\boxed{\hat{p} = 0.156}$$

$$(ii) P(x=0) = {}^4C_0 p^0 (1-p)^4$$

$$= 91.34$$

$$P(x=1) = {}^4C_1 (p)^1 (1-p)^3$$

$$= 67.53$$

$$P(x=2) = {}^4C_2 (p)^2 (1-p)^2$$

$$= 18.72$$

$$P(x=3) = {}^4C_3 (p)^3 (1-p)^1$$

$$= 2.31$$

$$P(x=4) = {}^4C_4 (p)^4 (1-p)^0$$

$$= 0.11$$

$h_i, f(x_i)$	O_i	E_i	$\frac{(O_i - E_i)^2}{E_i}$
$P(x=0)$	86	91.34	0.3123
$P(x=1)$	75	67.53	6.8263
$P(x=2)$	19	17.72	0.3982
$P(x=3)$	2	2.31	0.0416
$P(x=4)$	1	0.11	7.2009
			<u>38.8503</u>

$$\chi^2_{cal} = 38.8503 \text{ at } df = 3$$

$$\chi^2_{tab} = 7.815$$

$$\chi^2_{cal} = 39.856$$

$$\chi^2_{tab} = 3.841$$

$$\chi^2_{cal} > \chi^2_{tab}$$

Thus we will reject H_0 at 5%.

$$(4) f(x) = \frac{cB^2}{(x+B)^4} : x > 0, B > 0$$

$$(i) \int_0^{\infty} f(x) dx = 1$$

$$cB^3 = \int_0^{\infty} \frac{1}{(x+B)^4} dx = 1$$

$$cB^3 = \left[\frac{(x+B)^{-3}}{-3} \right]_0^{\infty} = 1$$

$$\frac{cB^3}{B^3} = 1$$

$$\boxed{c = 3}$$

$$E(x) = \int_0^{\infty} x \cdot f(x) \cdot dx$$

$$= 3B^3 \int_0^{\infty} \frac{x}{(x+B)^4} dx$$

$$= 3B^3 \left[\frac{(x+B)^{-2}}{-2} - \frac{B(x+B)^{-3}}{-3} \right]_0^{\infty}$$

$$= 3B^3 \left[0 - 0 - \left(\frac{B^{-2}}{-2} + \frac{B^{-2}}{3} \right) \right]$$

$$= 3B^2 \left[\frac{1}{2B^2} + \frac{1}{3B^2} \right] = \frac{B}{2}$$

$$E(x^2) = \int_0^{\infty} x^2 \cdot f(x) \cdot dx = 3B^3 \int_0^{\infty} \frac{x^2}{(x+B)^4} dx$$

$$(ii) B = x, x_2, x_3, \dots, x_n$$

$$F(x) = \bar{y}$$

$$\boxed{\hat{B} = 2\bar{x}_n}$$

$$E(\hat{\beta}) = E(2\bar{X}_n) = \frac{2n}{n} \times \frac{B}{2} = B \quad \therefore \text{unbiased}$$

$$\begin{aligned} \text{(ii)} \quad \text{MSE} &= \text{Var}(b\bar{X}_n) + (\text{Bias}[b\bar{X}_n])^2 \\ &= b^2 \times \frac{3}{n} \times \frac{B^2}{n} + (E(b\bar{X}_n) - B)^2 \\ &= \frac{B^2}{n} \left[\frac{3b^2}{n} + n \left(\frac{b}{2} - 1 \right)^2 \right] \end{aligned}$$

(iv) It can be seen that $\hat{\beta} = 2\bar{X}_n$ is an unbiased estimator of B .
Estimate $b\bar{X}_n$ in (ii) where $b = \frac{2}{1+3/n}$ is a more biased estimator.

It is also consistent.

$$\textcircled{5} \text{(i)} \quad X \sim U(a+b)$$

$$\bar{X} = \frac{1}{2}(a+b)$$

$$S = \frac{1}{\sqrt{12}}(b-a)$$

$$\sqrt{12} + a = b$$

$$\bar{X} = \frac{1}{2}(a + \sqrt{12}S + a)$$

$$\bar{X} = \frac{1}{2}[2a + \sqrt{12}S]$$

$$\hat{a} = \bar{X} - \sqrt{3}S$$

$$(ii) \bar{x} = \frac{1}{2}(a+b)$$

$$s = \frac{1}{\sqrt{2}}(b-a)$$

$$a = b - \sqrt{2}s$$

$$\bar{x} = \frac{1}{2}[b - \sqrt{2}s + b]$$

$$= \frac{1}{2}[2b - \sqrt{2}s]$$

$$\boxed{\hat{b} = \bar{x} + \sqrt{3}s}$$

(iii) Sample: 1, 2, 3, 4, 10.

$$\bar{x} = \frac{20}{5} = 4$$

$$s^2 = \frac{1}{5} \times 81 = \frac{27}{4}$$

$$\hat{\sigma} = \bar{x} - \sqrt{3} \times \frac{27}{4}$$

$$= 4 - \sqrt{3} \times \frac{27}{4} = -7.6913$$

$$\hat{b} = 4 + \sqrt{3} \left(\frac{27}{4} \right) = 15.69134, \text{ hence the result is false.}$$

$$(iv) f(n) = \frac{1}{b-a} = \frac{1}{b}$$

$$hL = -n \log b$$

$$\frac{dh}{db} = -\frac{n}{b} = 0 \rightarrow b \rightarrow \infty$$

$$\frac{d^2h}{db^2} = \frac{n}{b^2} > 0, \text{ minimum}$$

$$\boxed{\hat{\sigma}_b = \max. n}$$

(6) $H_0: p = 0.1$

$H_1: p > 0.1$

(i) $P(\text{type I}) = P[\text{Reject } H_0 | H_0 \text{ is T}]$
 $= P[X \geq 3 | p = 0.1] + P[X=0 | p=0.1] \cdot P[Y \geq 5 | p=0.1]$
 $+ P[X=1 | p=0.1] \cdot P[Y \geq 4 | p=0.1] + P[X=2 | p=0.1] \cdot P[Y \geq 3 | p=0.1]$
 $= 0.14727 \approx 15\%$

(ii) $P(X \geq 3 | p=0.3) + P(X=0 | p=0.3) = 0.91253 \approx 91\%$

(iii) $P(\text{type II error}) = 1 - 0.91253 = 0.08747 \approx 9\%$

(iv) $X \sim \text{Bin}(120, 1/3)$, $\beta = \frac{40}{120} = 1/3$

$$\frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}} \sim N(0,1)$$

$$P\left[-1.2816 < \frac{\hat{p} - p}{\sqrt{\frac{pq}{n}}} < 1.2816\right]$$

99 CI = (0.333, 0.2782)

(v) $\frac{X - np_0}{\sqrt{np_0q_0}} \sim N(0,1) \approx 1.2511$

(vii) $P.d = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$

$S_p^2 = 40270$

≤ 0.2034

(7) (i) Public =

$$\bar{x}_1 = 30$$

$$s^2 = 8.0195 \quad \sigma^2 = 64.3124$$

Private:

$$\bar{x}_2 = 35.06$$

$$s^2 = 8.2099 \quad \sigma^2 = 67.4024$$

(ii) $H_0: \bar{x}_1 = \bar{x}_2$ $H_1: \bar{x}_1 \neq \bar{x}_2$

$$\frac{(30 - 35.06)}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{(30 - 35.06)}{\sqrt{8.1133} \sqrt{\frac{2}{9}}} = -1.3226$$

$$s_p^2 = \frac{8(64.3124) + 8(67.4024)}{16} = 65.8579$$

$$Z_{\text{fail}} = (-2.624, 2.624)$$

~~accept~~ H_0 , not reject H_0 .

(iii) Private hospital don't result in higher claim.

(8)

(i) $E(\hat{\theta}) = \theta$, it will be unbiased.

(ii) $E(\hat{\theta}) \neq \theta$, then will be biased of parameter

(iii) $MSE = \text{var} + \text{bias}^2$.

- (iv) $\hat{\theta}$ is ^{more} efficient than $\tilde{\theta}$, since MSE of $\hat{\theta}$ is lower than $\tilde{\theta}$.
- (v) when MSE more toward 0, as n increase, the estimate would be more consistent.
- (vi) Method of Moments.
Maximum likelihood estimates.

(a) $t_k \sim \frac{Z}{\sqrt{\frac{\chi_k^2}{k}}}$

$Z = N(0,1)$
 χ_k^2 = Chi-square dist with degree of freedom k
 k = degree of freedom.

(i) $\mu(t_k) = 0$ $\text{Var}(t_k) = \frac{k}{k-2}$

(iii) $n = 10$
 $\mu = 50$
 $\sigma^2 = 48.667$

$t_{0.05, 9} = 3.25$

(a) $CI = 50 \pm 3.25 \sqrt{\frac{48.667}{10}} = (42.823, 57.1699)$

(b) $CI = 50 \pm 2.5758 \sqrt{\frac{48.667}{10}} = (44.3176, 55.6823)$

(10) $n=1000$, $X \sim \text{Pois}(3)$ $X \sim N(3000, 3000)$

(i) Type 1 error = $P[\text{Ho rejected} | \text{Ho is true}]$
 $= P[X > 3100 | \lambda = 3]$
 $= P\left[Z > \frac{3100 - 3000}{\sqrt{3000}}\right]$
 $= P[Z > 1.8257]$
 $= 1 - P[Z < 1.8257]$
 $= 3.89\%$

(ii) Type 2 error = $P[\text{do not reject Ho} | \text{Ho is false}]$
 $= P[X < 3100 | \lambda = n\lambda]$
 $= P\left[X < \frac{3100 - n\lambda}{\sqrt{n\lambda}}\right]$

(iii) Power = $P[\text{reject Ho} | \text{Ho is false}]$
 $= P[X > 3100 | \lambda = n\lambda]$
 $= 1 - P\left[\frac{3100 - n\lambda}{\sqrt{n\lambda}}\right]$

(iv) $n = 2900$.

$$CI = 3 \pm 2.5758 \sqrt{\frac{3}{2900}}$$

$$(2.9172, 3.0828)$$

$$\begin{aligned} \textcircled{11} \quad \Sigma x &= 213,200 \\ \Sigma x^2 &= 4,919,860,000 \\ \bar{x} &= 17,767 \\ S.d &= 10,144 \\ n &= 12 \end{aligned}$$

$$\Sigma x^1 = 213200 - 11000 - 48000 = 154200$$

$$\Sigma x^{11} = 154200 + 27500 + 31000 = 213200$$

$$\bar{x}^1 = 17767$$

$$\Sigma x^{21} = 4919860000 - 11000^2 - 48000^2 = 2494860000$$

$$\Sigma x^{211} = 2494860000 + 27500^2 + 31500^2 = 4243360000$$

$$S = \frac{1}{\sqrt{11}} \sqrt{4243360000 - 12(17767)^2}$$

$$S = 6434.0326$$

$$\textcircled{12} \quad X_1, X_2, X_3, \dots, X_n \sim U(0, \theta), \quad a=0, \quad b=\theta$$

$$\textcircled{1} \quad f(y) = \frac{\theta}{2}, \quad v(y) = \frac{\theta^2}{12}$$

$$L = \prod_{i=1}^n \left(\frac{1}{\theta} \right) = \frac{1}{\theta^n}$$

$$\ln(L) = -\frac{n}{\theta}$$

$$\frac{d \ln L}{d \theta} = -\frac{n}{\theta} \quad \lim_{\theta \rightarrow \infty} \theta = \text{infinite}$$

hence CRIB would not be applicable.

$$(ii) P[\text{Max } X < y] = \prod_{i=1}^n \left[\int_0^y \frac{1}{\theta} dx \right] = \left(\frac{y}{\theta} \right)^n$$

$$\begin{aligned} E[y^k] &= \int y^k \cdot \frac{n}{\theta^n} y^{n-1} dy \\ &= \frac{n}{n+k} \int_0^{\theta} \frac{y^{n+k}}{\theta^n} dy \\ &= \frac{n}{n+k} \cdot \left(\frac{\theta^{n+k} - 0}{\theta^n} \right) \\ &= \frac{n\theta^k}{n+k} \end{aligned}$$

$$\begin{aligned} (iv) \text{Bias}[\hat{\theta}(y)] &= E[\hat{\theta}(y) - \theta] \\ &= E[y - \theta] \\ &= cE[y] - \theta \\ &= c \cdot \frac{n\theta}{n+1} - \theta \\ &= \theta \left(\frac{cn}{n+1} - 1 \right) \end{aligned}$$

$$(iv) \frac{d}{dc} \text{MSE}[\hat{\theta}(y)] = \frac{d}{dc} \left[\frac{cn^2\theta^2}{n+2} - \frac{2n\theta^2}{n+1} (c\theta^2) \right]$$

$$c_m = \frac{2n\theta^2}{n+1} \cdot \frac{n+2}{2n\theta} = \frac{n+2}{n+1}$$

$$\frac{d^2}{dc^2} [\text{MSE}(\hat{\theta}(y))] = \frac{2n\theta^2}{n^2} > 0, \text{ min value}$$

(vi) $\frac{n+2}{n+1} y \rightarrow y$, two estimators become the same $n \rightarrow \infty$.