

Probability and Statistics Assignment - 2

1. Let X be the amount of a home insurance claim and Y the amount of a car insurance claim.

$$X \sim N(800, 100^2) \quad Y \sim N(1200, 300^2)$$

We require

$$P[(Y_1 + Y_2 + Y_3) > (X_1 + X_2 + X_3 + X_4) + 800]$$

$$= P[(Y_1 + Y_2 + Y_3) - (X_1 + X_2 + X_3 + X_4) > 800]$$

We need distribution of $(Y_1 + Y_2 + Y_3) -$

$(X_1 + X_2 + X_3 + X_4)$

$$= \sim N(400, 31000)$$

$$= P\left(Z > \frac{800 - 400}{\sqrt{31000}}\right) = P(Z > 0.8718)$$

$$= 1 - P(Z < 0.718)$$

$$= 0.236$$

3 (i) $f(x) = \lambda e^{-\lambda(x-\alpha)}, x \geq \alpha, \lambda > 0$

using MGF distribution

$$M_X(t) = E(e^{tx}) = \int_{\alpha}^{\infty} e^{tx} \lambda e^{-\lambda(x-\alpha)} \cdot dx$$

$$= \int_{\alpha}^{\infty} \lambda e^{\lambda\alpha} e^{-(\lambda-t)x} \cdot dx$$

$$= \left[\frac{\lambda e^{\lambda\alpha}}{\lambda-t} \cdot e^{-(\lambda-t)x} \right]_{\alpha}^{\infty}$$

$$= \frac{\lambda}{\lambda - t} \cdot e^{t\alpha} \quad \text{provided } t < \lambda$$

$$M_X(t) = \left(1 - \frac{t}{\lambda}\right)^{-1} e^{t\alpha}$$

$$M'_X(t) = \frac{1}{\lambda} \left(1 - \frac{t}{\lambda}\right)^{-2} \cdot e^{t\alpha} + \alpha \left(1 - \frac{t}{\lambda}\right)^{-1} \cdot e^{t\alpha}$$

$$E(X) = M'_X(0) = \frac{1}{\lambda} + \alpha$$

$$M''_X(t) = \frac{2}{\lambda^2} \left(1 - \frac{t}{\lambda}\right)^{-3} e^{t\alpha} + \frac{2\alpha}{\lambda} \left(1 - \frac{t}{\lambda}\right)^{-2} e^{t\alpha} + \alpha^2 \left(1 - \frac{t}{\lambda}\right)^{-1} e^{t\alpha}$$

$$E(X^2) = M''_X(0) = \frac{2}{\lambda^2} + \frac{2}{\lambda} \alpha + \alpha^2$$

$$\text{Var}(X) = \frac{2}{\lambda^2} + \frac{2}{\lambda} \alpha + \alpha^2 - \left(\frac{1}{\lambda} + \alpha\right)^2$$

$$\text{Var}(X) = \frac{1}{\lambda^2}$$

Prob, $P(X=n) = \frac{1}{n!} G_X^{(n)}(0)$

$$P(V=4) = \frac{1}{4!} G_V^{(4)}(0)$$

$$G'_V(t) = R p^k q (1 - qx)^{-k-1}$$

4.

$$f(x, y) = ax(x+y), \quad 0 < x < 5$$

we know that, $10 < y < 20$

$$\int_{10}^{20} \int_0^5 ax(x+y) dx dy = 1$$

$$\int_{10}^{20} \int_0^5 a(x^2 + xy) dx dy = 1$$

$$\int_{10}^{20} a \left[\frac{x^3}{3} + \frac{x^2 y}{2} \right]_0^5 dy = 1$$

$$\int_{10}^{20} a \cdot \left[\frac{125}{3} + \frac{25}{2} y \right] dy = 1$$

$$a \left[\frac{125}{3} y + \frac{25}{4} y^2 \right]_{10}^{20} = 1$$

$$a \cdot \left[\frac{2500}{3} + 2500 - \frac{1250}{3} - 625 \right] = 1$$

$$a(2291.666) = 1$$

$$a = 0.00044$$

$$E(Y/x) = \int_y y \cdot \frac{f(x, y)}{f(x)} \cdot dy$$

$$f(x) = \int_y f(x, y) \cdot dy$$

$$= \int_{10}^{20} a \cdot (x^2 + xy) dy$$

$$= a \cdot \int x^2 y + 2xy^2 \Big|_{10}^{20}$$

$$= a [20x^2 + 200x - 10x^2 - 50x]$$

$$= a [10x^2 + 150x]$$

$$E(Y/X) = \int_{10}^{20} y \frac{a[x^2 + 2y]}{a[10x^2 + 150x]} dy$$

$$= \int_{10}^{20} y \cdot \frac{(x+y)}{(10x+150)} dy$$

$$= \frac{1}{10x+150} \left[\frac{y^2}{2} x + \frac{y^3}{3} \right]_{10}^{20}$$

$$= \frac{1}{10x+150} \left(200x + \frac{8000}{3} - 50x - \frac{1000}{3} \right)$$

$$= \frac{1}{10x+150} \left(150x + \frac{7000}{3} \right)$$

Examining Arthur's Statements :

$$X \sim \text{Poi}(\lambda) \text{ \& \& } Y \sim \text{Poi}(\mu)$$

X has MGF

$$M_X(t) = e^{\lambda(e^t - 1)}$$

+

Y has MGF

$$M_Y(t) = e^{\mu(e^t - 1)}$$

so the difference is,

Page No.
 Date

$(X+Y)$ has MGF

$$\begin{aligned} M_X(t) \cdot M_Y(t) &= e^{\lambda(et-1)} \cdot e^{-\mu(et-1)} \\ &= e^{(\lambda-\mu)(et-1)} \end{aligned}$$

which is the MGF of a Poisson with parameters $\lambda - \mu$

So Arthur's statement is true

similarly,

$$X \sim \text{Poi}(\lambda) \text{ \& \& } Y \sim \text{Poi}(\mu)$$
$$M_X(t) = e^{\lambda(et-1)} \text{ \& \& } M_Y(t) = e^{\mu(et-1)}$$

so the sum $X+Y$ has MGF,

$$\begin{aligned} M_X(t) M_Y(t) &= e^{\lambda(et-1)} \cdot e^{\mu(et-1)} \\ &= e^{(\lambda+\mu)(et-1)} \end{aligned}$$

→ MGF of Poisson with parameter $\lambda + \mu$,

→ $X+Y \sim \text{Poi}(\lambda + \mu)$

∴ They're statement is also True.

$$(i) \text{Var}(X/Y=2) = E(X^2/Y=2) - E^2(X/Y=2)$$

$$E(X/Y=2) = 1 \cdot \frac{0}{0.6} + 2 \times \frac{0.3}{0.6} + 3 \times \frac{0.1}{0.6} + 4 \times \frac{0.2}{0.6}$$

$$= \frac{17}{6}$$

$$E(X^2/Y=2) = 1^2 \times \frac{0}{0.6} + 2^2 \times \frac{0.3}{0.6} + 3^2 \times \frac{0.1}{0.6} + 4^2 \times \frac{0.2}{0.6}$$

$$= \frac{53}{6}$$

$$\therefore \text{Var}(X/Y=2) = \frac{53}{6} - \left(\frac{17}{6}\right)^2$$

$$= \frac{29}{36} = 0.80556$$

Conditional Expectation

$$E(U/V=v) = \int u f(u/v) du$$

$$f(v) = \int_{u=0}^v \frac{48}{67} (2uv - u^2) du$$

$$= \frac{48}{67} \left[u^2 v - \frac{u^3}{3} \right]_0^v$$

$$= \frac{48}{67} \left[v - \frac{1}{3} \right]$$

$$\Rightarrow f(u/v) = \frac{f(u,v)}{f(v)} = \frac{\frac{48}{67} (2uv - v^2)}{\frac{48}{67} \left(v - \frac{1}{3} \right)}$$

$$= \frac{(2uv - v^2)}{\left(v - \frac{1}{3} \right)}$$

Page No.
 Date

$$\text{So, } E(u/v - v) = \int_0^1 \frac{2u^2 v - u^3}{v - \frac{1}{3}} du$$

$$= \left[\frac{2/3 u^3 v - \frac{u^4}{4}}{v - \frac{1}{3}} \right]_0^1$$

$$= \frac{\frac{2}{3} v - \frac{1}{4}}{v - \frac{1}{3}}$$

8. $X \sim \text{Gamma}(3, 2)$
 $E(Y/X = x) = 3x + 1$
 $\text{var}(Y/X = x) = 2x^2 + 5$

$$E(Y) = E(E(Y/X)) = E(3x + 1)$$

$$= 3E(X) + 1$$

but $E(x) = \frac{3}{2} = 1.5$

$$E(Y) = 3(1.5) + 1$$

$$= 4.5 + 1$$

$$E(Y) = 5.5$$

using $\text{var}(Y) = E(\text{var}(Y/X)) + \text{var}[E(Y/X)]$

$$= \text{var}(Y) = E(2x^2 + 5) + \text{var}(3x + 1)$$

$$= 2E(x^2) + 5 + 9\text{var}(x)$$

$$= 2(1.5)^2 + \left(\frac{9}{16}\right) + 9\left(\frac{3}{4}\right)$$

$$= 2(2.25) + 6.75$$

$$\text{var}(Y) = \underline{\underline{12.75}}$$

$$E(X) = 3$$

$$Sdx = 2$$

$$E(Y) = 4$$

$$Sdy = 1$$

$$\text{Correlation coefficient} = -0.3$$

$$Z = X + Y$$

$$\begin{aligned} \text{(i)} \quad \text{Cov}(X, Z) &= \text{Cov}(X, X + Y) \\ &= \text{Cov}(X, X) + \text{Cov}(X, Y) \\ &= \text{var}(X) + \text{Cov}(X, Y) \end{aligned}$$

using correlation coefficient between X & Y given:

$$\text{Corr}(X, Y) = -0.3 = \frac{\text{Cov}(X, Y)}{\sqrt{\text{var}(X) \text{var}(Y)}} = \frac{\text{Cov}(X, Y)}{\sqrt{4 \times 1}}$$

$$\text{Cov}(X, Y) = -0.6$$

$$\text{Cov}(X, Z) = 4 - 0.6 = 3.4$$

$$\text{(ii)} \quad \text{var}(Z) = \text{Cov}(Z, Z)$$

$$\begin{aligned} \text{var}(Z) &= \text{Cov}(X + Y, X + Y) \\ &= \text{Cov}(X, X) + 2\text{Cov}(X, Y) + \text{Cov}(Y, Y) \\ &= \text{var}(X) + 2\text{Cov}(X, Y) + \text{var}(Y) \\ &= 4 + 2(-0.6) \\ &= 3.8 \end{aligned}$$

10.

$$k x^{-\alpha} e^{-y/\beta}$$

$$1 < x < \infty$$

$$1 < y < \infty$$

since the pdf must integrate to 1

$$\int_{y=1}^{\infty} \int_{x=1}^{\infty} k x^{-\alpha} e^{-y/\beta} dx \cdot dy = 1$$

$$= \int_1^{\infty} k x^{-\alpha} e^{-y/\beta} = k e^{-y/\beta} \left[\frac{x^{-\alpha+1}}{-\alpha+1} \right]_1^{\infty}$$

$$= \int_1^{\infty} \frac{k e^{-y/\beta}}{\alpha-1} = \frac{k}{\alpha-1} \left[-\beta e^{-y/\beta} \right]_1^{\infty}$$

$$= \frac{k \cdot \beta e^{-1/\beta}}{\alpha-1}$$

Equating it to 1

$$\frac{k \beta e^{-1/\beta}}{\alpha-1} = 1$$

$$k = \frac{(\alpha-1) e^{1/\beta}}{\beta}$$