

PROBABILITY AND STATISTICS

ASSIGNMENT-2

- 1] Let X be the amount of home insurance claim
 Y be the amount of car insurance claim

$$X \sim N(800, 100^2) \quad Y \sim N(1200, 300^2)$$

We require

$$P(Y_1 + Y_2 + Y_3 > (X_1 + X_2 + X_3 + X_4) + 800)$$

$$= P(Y_1 + Y_2 + Y_3 - (X_1 + X_2 + X_3 + X_4) > 800)$$

We need distribution of $(Y_1 + Y_2 + Y_3) - (X_1 + X_2 + X_3 + X_4)$:

$$\sim N(3 \times 1200 - 4 \times 800, 3 \times 300^2 + 4 \times 100^2)$$

$$\sim N(400, 310000)$$

$$= P\left(Z > \frac{(360 - 400)}{\sqrt{310000}}\right)$$

$$P(Z > 0.8718)$$

$$1 - P(Z < 0.8718) = 0.230$$

3] i) $f(x) = \lambda e^{-\lambda(x-\alpha)}$

Using the definition of MGF:

$$M_X(t) = E[e^{tx}] = \int_{-\infty}^{\infty} e^{tx} \lambda e^{-\lambda(x-\alpha)} dx$$

$$= \int_{-\infty}^{\infty} \lambda e^{\lambda\alpha} e^{-(\lambda-t)x} dx$$

$$= \left[-\frac{\lambda e^{\lambda\alpha}}{\lambda-t} e^{-(\lambda-t)x} \right]_{-\infty}^{\infty}$$

$$= \frac{\lambda}{\lambda-t} e^{t\alpha}, \quad \text{given } t < \lambda$$

$$ii) M_x(t) = \left(1 - \frac{t}{\lambda}\right)^{-1} e^{tx}$$

$$M'_x(t) = \frac{1}{\lambda} \left(1 - \frac{t}{\lambda}\right)^{-1} e^{\alpha} + \alpha \left(1 - \frac{t}{\lambda}\right)^{-1} e^{tx}$$

$$E(x) = M'_x(0) = \frac{1}{\lambda} + \alpha$$

$$M''_x(t) = \frac{2}{\lambda^2} \left(1 - \frac{t}{\lambda}\right)^{-2} e^{tx} + \frac{2\alpha}{\lambda} \left(1 - \frac{t}{\lambda}\right)^{-1} + \alpha^2 \left(1 - \frac{t}{\lambda}\right)^{-1} e^{tx}$$

$$E(x^2) = M''_x(0) = \frac{2}{\lambda^2} + \frac{2\alpha}{\lambda} + \alpha^2$$

$$\text{Var}(x) = \frac{2}{\lambda^2} + \frac{2\alpha}{\lambda} + \alpha^2 - \left(\frac{1}{\lambda} + \alpha\right)^2$$

$$\text{Var}(x) = \frac{1}{\lambda^2}$$

$$4] \text{ Probability, } P(X=n) = \frac{1}{n!} G_v^{(n)}(0)$$

$$P(X=4) = \frac{1}{4!} G_v^{(4)}(0)$$

$$\text{finding } G_v^{(4)}(t):$$

$$G_v'(t) = k p^k q (1 - qe^t)^{-k-1}$$

$$G_v''(t) = -k(k-1) k p^k q^2 (1 - qe^t)^{-k-2}$$

$$G_v'''(t) = (-k-2)(-k-1) k p^k q^3 (1 - qe^t)^{-k-3}$$

$$G_v^{(4)}(t) = (-k-3)(-k-2)(-k-1) k p^k q^4 (1 - qe^t)^{-k-4}$$

$$G_v^{(4)}(0) = (-k-3)(-k-2)(-k-1) k p^k q^4$$

$$4. \quad f(r, y) = ar(r+y) \quad 0 < r < 5$$

We know that,

$$\int_{10}^{20} \int_0^5 ar(r+y) dr dy = 1$$

$$\int_{10}^{20} \int_0^5 a(r^2 + ry) dr dy = 1$$

$$\int_{10}^{20} a \left[\frac{r^3}{3} + \frac{r^2 y}{2} \right]_0^5 dy = 1$$

$$\int_{10}^{20} a \left[\frac{125}{3} + \frac{25}{2} y \right] dy = 1$$

$$a \left[\frac{125y}{3} + \frac{25y^2}{4} \right]_{10}^{20} = 1$$

$$a \left[\frac{2500}{3} + 2500 - \frac{1250}{3} - 625 \right] = 1$$

$$a = 0.00044$$

$$5] \quad E(Y|X) = \int_y \frac{y f(r, y)}{f(r)} dy$$

$$f(r) = \int_y f(r, y) dy$$

$$= \int_{10}^{20} a(r^2 + ry) dy$$

$$= a \left[r^2 y + \frac{ry^2}{2} \right]_{10}^{20}$$

$$= a [20r^2 + 260r - 10r^2 - 50r]$$

$$= a [10r^2 + 150r]$$

$$E(Y/X) = \int_{10}^{20} y \frac{a[x^2 + xy]}{a[10x^2 + 150x]} dy$$

$$\int_{10}^{20} y \frac{(x+y)}{(10x+150)} dy$$

$$= \frac{1}{10x+150} \left[\frac{y^2 x}{2} + \frac{y^3}{3} \right]_{10}^{20}$$

$$= \frac{1}{10x+150} \left(\frac{200x + 3000}{3} - \frac{50x + 1000}{3} \right)$$

$$= \frac{1}{10x+150} (150x + 2000)$$

d) Examining Arthur's statement,

$$X \sim \text{Poi}(\lambda) \quad Y \sim \text{Poi}(\mu)$$

$$X \text{ has MGF: } M_X(t) = e^{\lambda(e^t - 1)}$$

$$Y \text{ has MGF: } M_Y(t) = e^{\mu(e^t - 1)}$$

$$\begin{aligned} X-Y \text{ has MGF: } M_{X-Y}(t) &= M_X(t) M_Y(-t) \\ &= e^{\lambda(e^t - 1)} e^{-\mu(e^t - 1)} \\ &= e^{(\lambda - \mu)(e^t - 1)} \end{aligned}$$

It is the MGF of a Poisson with parameter $(\lambda - \mu)$
 $X - Y \sim \text{Poi}(\lambda - \mu)$

So Arthur's statement is true

Similarly evaluating Christine's statement

$$X \sim \text{Poi}(\lambda)$$

$$Y \sim \text{Poi}(\mu)$$

$$M_X(t) = e^{\lambda(e^t - 1)}$$

$$M_Y(t) = e^{\mu(e^t - 1)}$$

So $X+Y$ has MGF: $M_X(t) M_Y(t)$
 $= e^{\lambda(e^t-1)} e^{\mu(e^t-1)}$
 $= e^{(\lambda+\mu)(e^t-1)}$

It is @ MGF of a poisson with parameters $(\lambda+\mu)$
 $X+Y \sim \text{Poi}(\lambda+\mu)$

So Christine's Statement is also true

8) $X \sim \text{Gamma}(3, 2)$

$$E(Y|X=x) = 3x+1$$

$$\text{Var}(Y|X=x) = 2x^2+5$$

$$E(Y) = E(E(Y|X)) = E(3X+1)$$

$$= 3E(X) + 1$$

$$\text{But } E(X) = \frac{3}{2}$$

$$E(Y) = 3(1.5) + 1$$

$$= 5.5$$

$$\text{Using } \text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X))$$

$$= \text{Var}(Y) = E(2x^2+5) + \text{Var}(3x+1)$$

$$= 2E(x^2) + 5 + 9\text{Var}(X)$$

$$= 2\left((1.5)^2 + \frac{9}{16}\right) + 9\left(\frac{3}{4}\right)$$

$$\text{Var}(Y) = 12.375$$

$$9] E[X] = 3$$

$$\sigma_x = 2$$

$$E[Y] = 4$$

$$\sigma_y = 1$$

$$\text{Correlation coefficient} = -0.3$$

$$Z = X + Y$$

$$\begin{aligned} i) \text{Cov}(X, Z) &= \text{Cov}(X, X+Y) \\ &= \text{Cov}(X, X) + \text{Cov}(X, Y) \\ &= \text{Var}(X) + \text{Cov}(X, Y) \end{aligned}$$

$$\text{Cov}(X, Y) = -0.3 = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \cdot \text{Var}(Y)}} = \frac{\text{Cov}(X, Y)}{\sqrt{4 \times 1}}$$

$$\text{Cov}(X, Y) = -0.6$$

$$\text{Cov}(X, Z) = 4 - 0.6 = 3.4$$

$$\begin{aligned} ii) \text{Var}(Z) &= \text{Cov}(Z, Z) \\ &= \text{Cov}(X+Y, X+Y) \\ &= \text{Cov}(X, X) + 2\text{Cov}(X, Y) + \text{Cov}(Y, Y) \\ &= \text{Var}(X) + 2\text{Cov}(X, Y) + \text{Var}(Y) \\ &= 4 + 2(-0.6) + 1 \\ &= 3.8 \end{aligned}$$

$$10] k x^{-\alpha} e^{y/3} \quad 1 < x < \infty$$

The PDF must integrate to 1

$$\int_{y=1}^{\infty} \int_{x=1}^{\infty} k x^{-\alpha} e^{y/3} dx dy = 1$$

Integrating over x values given

$$\int_1^{\infty} k x^{-\alpha} e^{-x/\beta} = k e^{-1/\beta} \left[\frac{x^{-\alpha+1}}{-\alpha+1} \right]_1^{\infty}$$

$$= \frac{k e^{-1/\beta}}{\alpha-1}$$

Integrating over y values

$$\int_1^{\infty} \frac{k e^{-y/\beta}}{\alpha-1} = \frac{k}{\alpha-1} \left[-\beta e^{-y/\beta} \right]_1^{\infty}$$

$$= \frac{k \beta e^{-1/\beta}}{\alpha-1}$$

Equating to 1

$$\frac{k \beta e^{-1/\beta}}{\alpha-1} = 1$$

$$k = \frac{(\alpha-1) e^{1/\beta}}{\beta}$$